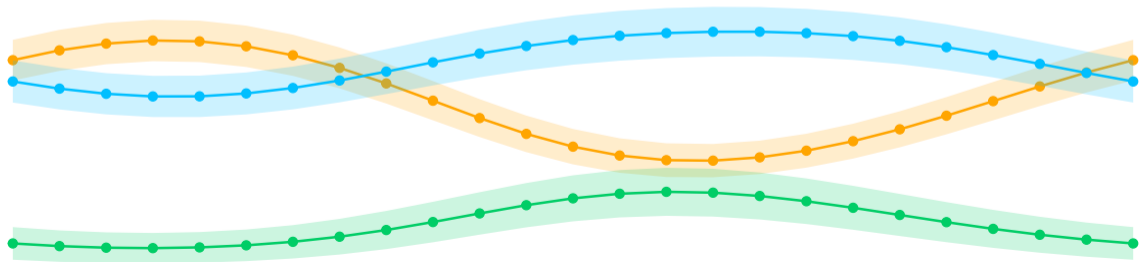


Young Researchers Workshop

Inference on the state process of periodically inhomogeneous hidden Markov models

Jan-Ole Koslik

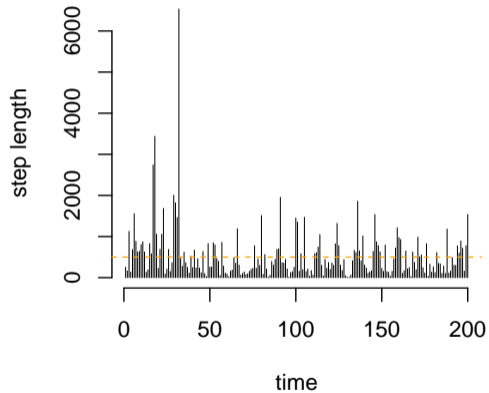
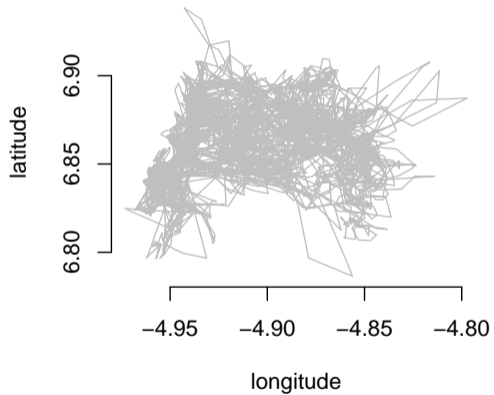
February 16, 2024



Motivating example

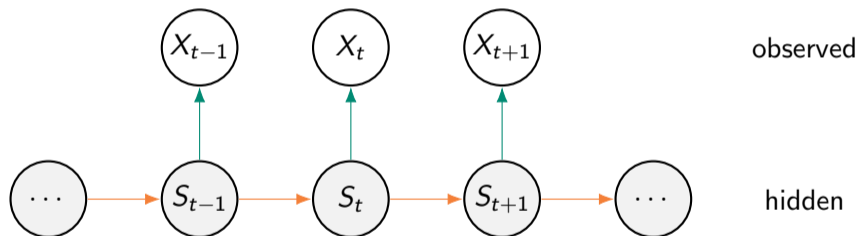


Motivating example



Quick Recap on HMMs

Doubly stochastic process:



- ▶ every observation is generated by one of N possible distributions f_1, \dots, f_N ,
- ▶ the state process selects which distribution is active at any given time point

More formal definition of an HMM

An N -state HMM is a (doubly) stochastic process in discrete time, with

- ▶ an unobserved **state process** S_1, S_2, \dots, S_T taking values in $\{1, \dots, N\}$,
- ▶ and an **observation process** X_1, X_2, \dots, X_T ,

such that

- ▶ $f(s_t | s_1, \dots, s_{t-1}) = f(s_t | s_{t-1})$
(Markov property)
- ▶ $f(x_t | s_1, \dots, s_T, x_1, \dots, x_{t-1}, x_{t+1}, \dots, x_T) = f(x_t | s_t)$
(conditional independence assumption)

Reminder: Markov chains

The **Markovian state process** is fully characterised by the initial distribution

$$\delta^{(1)} = (\Pr(S_1 = 1), \dots, \Pr(S_1 = N))$$

and the **transition probabilities**

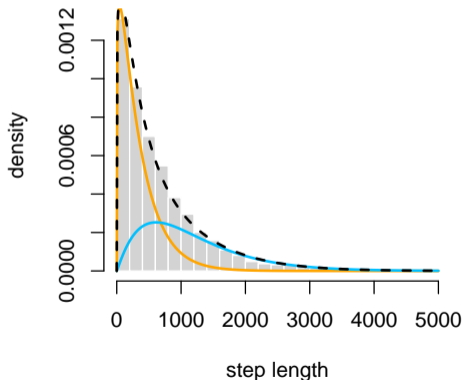
$$\gamma_{ij}^{(t)} = \Pr(S_{t+1} = j \mid S_t = i).$$

which we summarise in the **transition probability matrix** (t.p.m.)

$$\mathbf{\Gamma}^{(t)} = (\gamma_{ij}^{(t)})_{i,j=1,\dots,N}.$$

Example continued

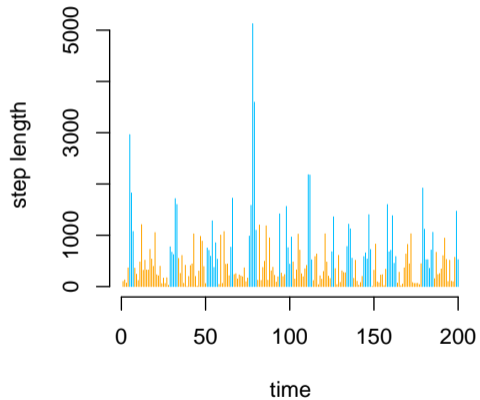
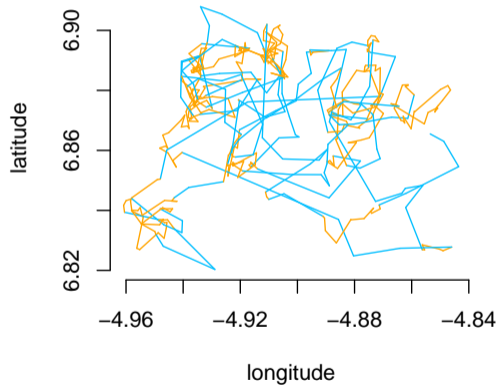
Fitting a basic 2-state HMM to the elephant data yields the following results:



$$\hat{\mathbf{\Gamma}} = \begin{pmatrix} 0.81 & 0.19 \\ 0.29 & 0.71 \end{pmatrix}$$

$$\hat{\boldsymbol{\delta}} = (0.6, 0.4)$$

Example continued



Example continued (the problems)

... which is all very nice, **but**

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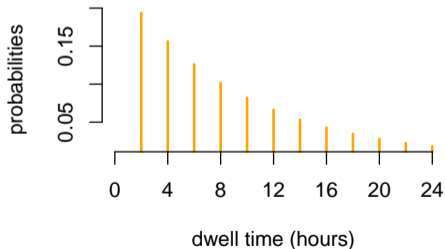
1. our model assumes the unconditional state probabilities are constant,

Example continued (the problems)

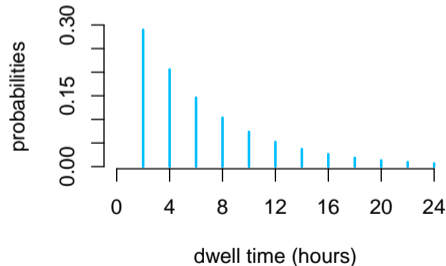
... which is all very nice, **but**

1. our model assumes the unconditional state probabilities are constant,
2. and the distribution of time spent in a state is geometric (often criticised)
(e.g. for state 1: $\Pr(R_1 = r) = 0.81^{r-1}0.19$)

inactive state



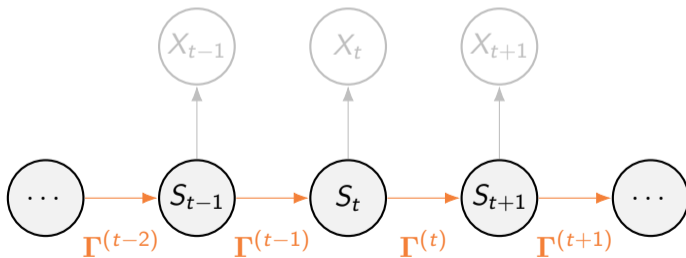
active state



Periodic variation

In the real world, there are many reasons for processes to involve a periodic component. For animal data, this could be

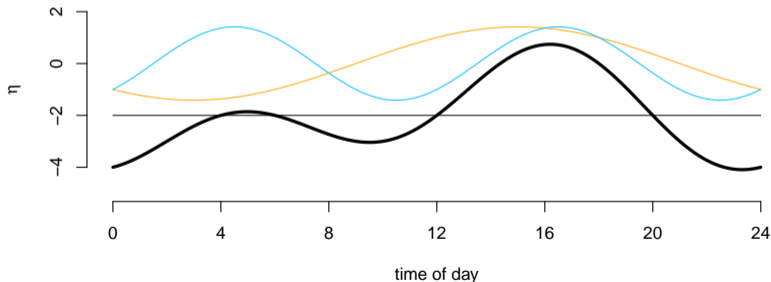
- ▶ diurnal/ time-of-day variation,
- ▶ seasonal variation,
- ▶ migratory behaviour.



How to model periodic variation

Typical parametric¹ model:

$$\gamma_{ij}^{(t)} = \text{mlogit}^{-1} \left(\beta_0^{(ij)} + \sum_{k=1}^K \beta_{1k}^{(ij)} \sin\left(\frac{2\pi kt}{L}\right) + \sum_{k=1}^K \beta_{2k}^{(ij)} \cos\left(\frac{2\pi kt}{L}\right) \right)$$



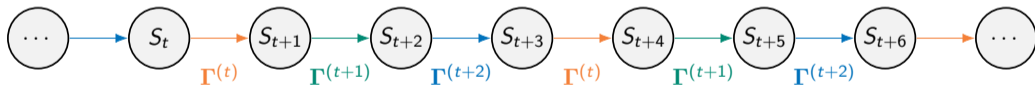
¹We can also do this non-parametrically (Feldmann et al., 2023).

Periodic variation

For a cycle length of L , both options lead to

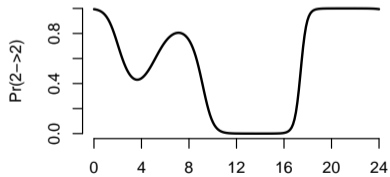
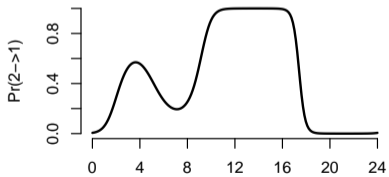
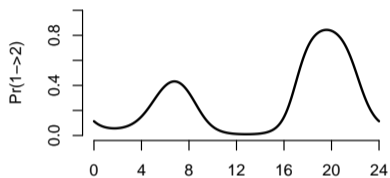
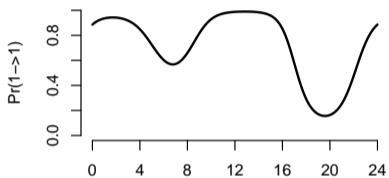
$$\Gamma^{(t)} = \Gamma^{(t+L)} \quad \text{for all } t = 1, \dots, T, \quad (1)$$

which is what we mean by **periodic variation**.



Bad interpretability

$\hat{\Gamma} =$



Periodic stationarity

- ▶ We would like to have $\Pr(S_t = i)$ as a function of the time of day.
- ▶ For periodic variation, this is typically approximated² by $\boldsymbol{\rho}^{(t)}$ which is the solution to

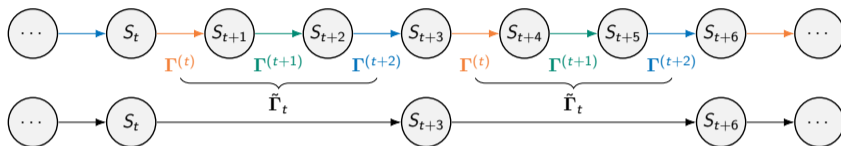
$$\boldsymbol{\rho}^{(t)}\boldsymbol{\Gamma}^{(t)} = \boldsymbol{\rho}^{(t)}, \quad \text{s.t.} \quad \sum_{i=1}^N \rho_i^{(t)} = 1, \quad \text{for } t = 1, \dots, L.$$

- ▶ But this estimate is biased, as it ignores the preceding process dynamics.

²e.g. Patterson et al., 2009

Periodic stationarity

- ▶ Consider for every $t \in \{1, \dots, L\}$ the thinned Markov chain $S_t, S_{t+L}, S_{t+2L}, \dots$

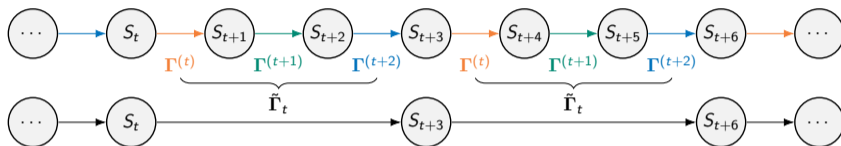


- ▶ It has constant t.p.m.

$$\tilde{\Gamma}_t = \Gamma^{(t)} \cdot \Gamma^{(t+1)} \cdot \dots \cdot \Gamma^{(t+L-1)}.$$

Periodic stationarity

- Consider for every $t \in \{1, \dots, L\}$ the thinned Markov chain $S_t, S_{t+L}, S_{t+2L}, \dots$



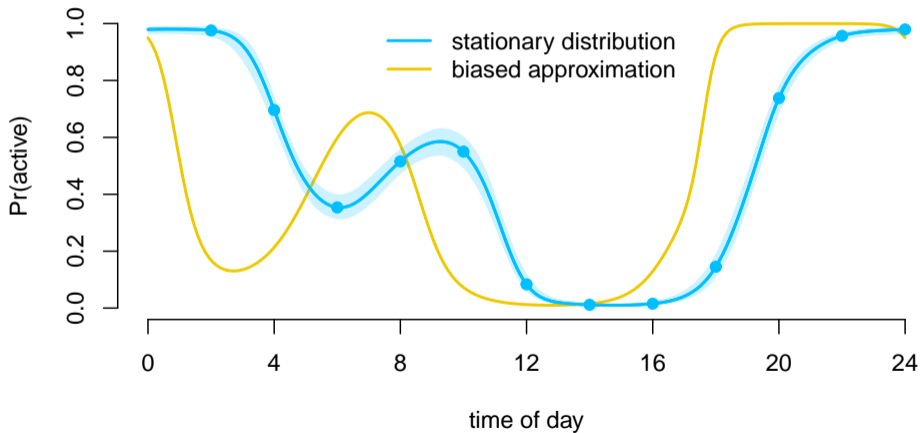
- It has constant t.p.m.

$$\tilde{\Gamma}_t = \Gamma^{(t)} \cdot \Gamma^{(t+1)} \cdot \dots \cdot \Gamma^{(t+L-1)}.$$

- Thus each thinned chain converges, and we get $\Pr(S_t = i)$ for each t by solving

$$\delta^{(t)} \tilde{\Gamma}_t = \delta^{(t)} \quad \text{s.t.} \quad \sum_{i=1}^N \delta_i^{(t)} = 1.$$

Periodically stationary distribution (Elephant example)



Dwell-time distribution(s)

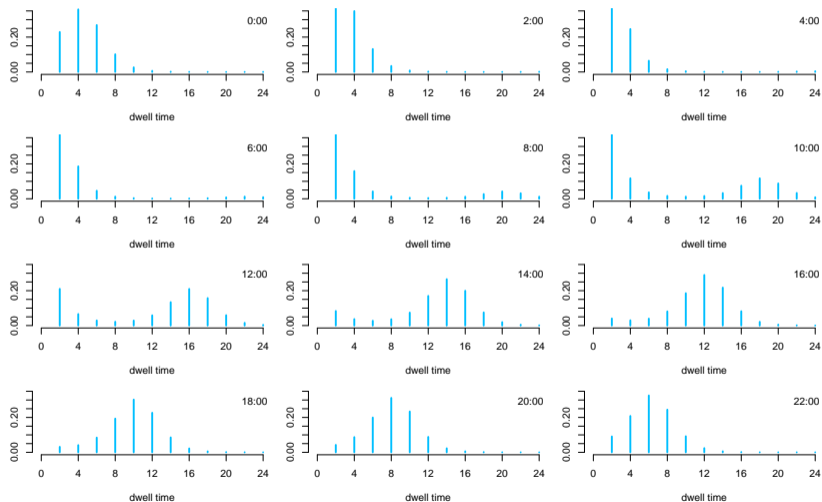
- ▶ We added periodic variation to our model, which is much more realistic.
- ▶ But what happened to the dwell-time distributions? Did it help?

Time-varying dwell-time distribution

First, consider the dwell-time distribution only at a certain time point t :

$$d_i^{(t)}(r) = \underbrace{(1 - \gamma_{ii}^{(t+r-1)})}_{\text{leave}} \cdot \underbrace{\prod_{j=1}^{r-1} \gamma_{ii}^{(t+j-1)}}_{\text{stay } r \text{ times}}, \quad r \in \mathbb{N}$$

Time-varying dwell-time distribution (Elephant example)



Time-varying dwell-time distribution

- ▶ The **time-varying** dwell-time distributions allow for detailed inference regarding the state process.
- ▶ However, it might be cumbersome to interpret all of them (e.g. minute-by-minute data: $L = 1440$).
- ▶ Focus of the inference is often the **overall** distribution of time spent in each state...

Overall dwell-time distribution

For a periodically stationary Markov chain defined by $\Gamma^{(t)}, t = 1, \dots, L$, the p.m.f. of the dwell-time distribution in state i is

$$d_i(r) = \sum_{t=1}^L w_i^{(t)} d_i^{(t)}(r), \quad r \in \mathbb{N}$$

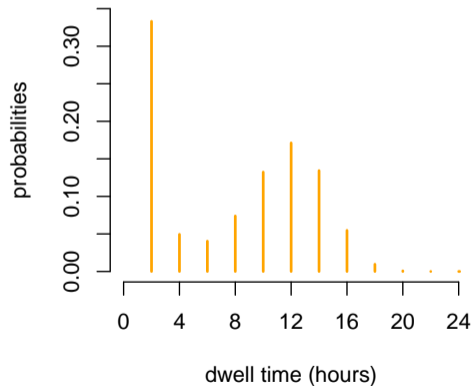
where with the mixture weights defined as

$$w_i^{(t)} = \frac{\sum_{l \neq i} \delta_l^{(t-1)} \gamma_{li}^{(t-1)}}{\sum_{t=1}^L \sum_{l \neq i} \delta_l^{(t-1)} \gamma_{li}^{(t-1)}}, \quad t = 1, \dots, L,$$

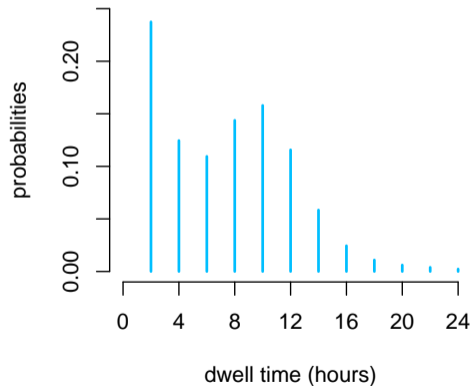
where $\delta^{(t)}$ is the periodically stationary distribution.

Overall dwell-time distribution (Elephant example)

inactive state



active state



Conclusion and Outlook

- ▶ In HMM including periodic variation, dwell-time distributions can be very non-geometric, i.e.
- ▶ it alleviates biologically unrealistic consequences of the Markov assumption.
- ▶ Thus, criticising the Markov assumption for implying geometric dwell times falls short of such models' actual potential.

Literature

- Feldmann, Carlina C et al. (2023). “Flexible modelling of diel and other periodic variation in hidden Markov models”.
In: *Journal of Statistical Theory and Practice*. in press.
- Patterson, Toby A. et al. (2009). “Classifying movement behaviour in relation to environmental conditions using hidden Markov models”. In: *Journal of Animal Ecology* 78.6, pp. 1113–1123.
- Wikelski M Davidson SC, Kays R. (2024). *Movebank: archive, analysis and sharing of animal movement data*.
Hosted by the Max Planck Institute of Animal Behavior. www.movebank.org (Accessed: 30.01.2024).

Thank you very much for your attention!

